# Optimal plot size in buckwheat 

# Tamanho ótimo de parcela em trigo mourisco 

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## Highlights

The three methods for determining the optimal plot size generate different results.
A plot size of $7.60 \mathrm{~m}^{2}$ is the reference for experiments with buckwheat. Information is provided for improving experimental precision in buckwheat.


#### Abstract

The aim of this work was to compare three methods of estimating the optimal plot size for evaluating fresh matter in the IPR91-Baili and IPR92-Altar cultivars of buckwheat (Fagopyrum esculentum Moench). Sixteen uniformity trials (blank experiments) were conducted, eight with the IPR91-Baili cultivar and eight with the IPR92-Altar cultivar. The trials were carried out at eight different sowing times. The fresh matter was evaluated in 576 basic experimental units (BEU), each $1 \mathrm{~m} \times 1 \mathrm{~m}$ in size ( 36 BEU per trial). The optimal plot size was determined using the method of modified maximum curvature, the linear response plateau model and the quadratic response plateau model. The optimal plot size differs between methods, and decreases in the following order: quadratic response plateau model, linear response plateau model and modified maximum curvature. The optimal plot size for evaluating fresh matter in the IPR91-Baili and IPR92-Altar cultivars of buckwheat is $7.60 \mathrm{~m}^{2}$. This size can be used as a reference for future experiments with buckwheat.


Key words: Soil cover crop. Experimental design. Uniformity trial. Fagopyrum esculentum Moench.

## Resumo

O objetivo deste trabalho foi comparar três métodos de estimação do tamanho ótimo de parcela para avaliar a massa de matéria fresca de trigo mourisco (Fagopyrum esculentum Moench) das cultivares IPR91Baili e IPR92-Altar. Foram conduzidos 16 ensaios de uniformidade (experimentos em branco), sendo oito com a cultivar IPR91-Baili e oito com a cultivar IPR92-Altar. Os ensaios foram realizados em oito épocas de semeadura. Foi avaliada a massa de matéria fresca em 576 unidades experimentais básicas (UEB) de $1 \mathrm{~m} \times$ 1 m (36 UEB por ensaio). Foi determinado o tamanho ótimo de parcela por meio dos métodos da curvatura

[^0]máxima modificado, do modelo linear de resposta com platô e do modelo quadrático de resposta com platô. O tamanho ótimo de parcela difere entre os métodos e decresce na seguinte ordem: modelo quadrático de resposta com platô, modelo linear de resposta com platô e curvatura máxima modificado. O tamanho ótimo de parcela para avaliar a massa de matéria fresca de trigo mourisco, das cultivares IPR91-Baili e IPR92Altar é de $7,60 \mathrm{~m}^{2}$. Esse tamanho pode ser utilizado como referência para futuros experimentos com trigo mourisco.
Palavras-chave: Cultura de cobertura de solo. Dimensionamento experimental. Ensaio de uniformidade. Fagopyrum esculentum Moench.

## Introduction

Buckwheat (Fagopyrum esculentum Moench) is an annual plant of family Polygonaceae, originally from Central Asia and grown in Europe, Asia and America (Accame \& Ortega, 2019). It displays fast growth, a short cycle, high tolerance to acidity and the ability to grow in low-fertility soils (Gonçalves et al., 2016). Important properties of buckwheat related to its use in human and animal nutrition, as a medicinal plant and as ground cover have been highlighted (Accame \& Ortega, 2019; Gonçalves et al., 2016; Görgen et al., 2016; Menezes \& Leandro, 2004; Mikhailovich, 2019; Pereira et al., 2017; Skora \& Campos, 2017; Yilmaz, Ayhan, \& Meriç, 2020; Ziech et al., 2015).

The buckwheat cultivars, IPR91-Baili and IPR92-Altar, are widely used experimental materials. It is important to know the variability of the soil and of the experimental material in order to correctly plan the plot size and choose the genetic base with the least variability. By using an appropriate plot size and least-variable genetic base, it is possible to minimise the experimental error and, consequently, increase the accuracy of the inferences. Smaller plot sizes require a smaller experimental area, but can generate less-reliable results due to the reduced representativeness of the experiment (sample). On the other hand, larger plot sizes can make it difficult to conduct the experiment
due to the necessary size of the experimental area, but would have the advantage of being more representative and, as a result, having greater experimental precision. Proper plot sizing optimises resources involved in the research, such as labour, time, financial resources and the experimental area.

Research in buckwheat together with other species of ground cover has been conducted with different plot sizes, such as: $25 \mathrm{~m}^{2}$ with a working area of $12 \mathrm{~m}^{2}$ (Menezes \& Leandro, 2004); $25 \mathrm{~m}^{2}$ (Ziech et al., 2015); $20 \mathrm{~m}^{2}$ (Görgen et al., 2016); $4 \mathrm{~m}^{2}$ (Pereira et al., 2017) and $24 \mathrm{~m}^{2}$ (Skora \& Campos, 2017). These surveys point out promising aspects of buckwheat and other species of ground cover, and no reference is made to defining the plot size for evaluating fresh matter. It is therefore important to determine the optimal plot size to be used as a reference in planning future experiments with buckwheat.

Among the various methods for estimating optimal plot size, the method of modified maximum curvature (Meier \& Lessman, 1971), the linear response plateau model (Paranaíba, Ferreira, \& Morais, 2009a) and the quadratic response plateau model (Peixoto, Faria, \& Morais, 2011) stand out. Comparative studies involving these methods and others have been carried out with rice (Paranaíba et al., 2009a), wheat and cassava (Paranaíba, Morais,
\& Ferreira, 2009b), passion fruit (Peixoto et al., 2011) and papaya (Brito, Faria, Morais, Sousa, \& Dantas, 2012), with different results between methods.

In uniformity trials (blank experiments), it is possible to divide the experimental area into basic experimental units (BEU) of the smallest possible size compatible with the evaluations (Storck, Garcia, Lopes, \& Estefanel, 2016). With the data collected in these BEUs, it is possible to form different plot sizes $(X)$ by grouping adjacent BEUs, and then estimate the coefficient of variation $\left(\mathrm{CV}_{(x)}\right)$ between the BEUs. The values of $\mathrm{CV}_{(x)}$ and X can be related using these three methods for determining the optimal plot size ( Xo ) and the coefficient of variation at the optimal plot size $\left(\mathrm{CV}_{\text {xo }_{0}}\right)$.

Carrying out uniformity trials for different sowing times and with different cultivars is important to increase the representativeness of the results, as it submits the crop to scenarios of environmental and genetic variability. Therefore, an analysis of this set of uniformity trials generates useful information to be used as a reference in planning experiments with buckwheat aiming at greater experimental precision. It is assumed that the optimal plot size differs between cultivars and between the methods of estimation. As such, the aim of this study was to compare three methods of estimating the optimal plot size for evaluating fresh matter in two cultivars (IPR91Baili and IPR92-Altar) of buckwheat [Fagopyrum esculentum (Moench)].

## Materials and Methods

Sixteen uniformity trials were conducted using buckwheat [Fagopyrum esculentum (Moench)], eight with the IPR91-

Baili cultivar and eight with the IPR92-Altar cultivar, in an experimental area located at $29^{\circ} 42^{\prime} \mathrm{S}$ and $53^{\circ} 49^{\prime} \mathrm{W}$, at an altitude of 95 m . According to the Köppen classification, the climate is type Cfa, humid subtropical, with hot summers and no dry season (Alvares, Stape, Sentelhas, Gonçalves, \& Sparovek, 2013); the soil is a Distrophic Arenic Red Argisol (Santos et al., 2018). A physical and chemical analysis of the soil at a depth of $0-20 \mathrm{~cm}$ revealed: $\mathrm{pH}_{\mathrm{H} 2 \mathrm{O}}$ 1:1:5.5, Ca: $4.7 \mathrm{cmol}_{\mathrm{c}} \mathrm{dm}^{-3}, \mathrm{Mg}: 1.9 \mathrm{cmol}_{\mathrm{c}} \mathrm{dm}^{-3}$, Al: $0.0 \mathrm{cmol}_{\mathrm{c}} \mathrm{dm}^{-3}, \mathrm{H}+\mathrm{Al}: 4.4 \mathrm{cmol}_{\mathrm{c}} \mathrm{dm}^{-3}, \mathrm{SMP}$ index: 6.0, organic matter: $2.3 \%$, clay content: 29.0\%, S: $1.5 \mathrm{mg} \mathrm{dm}^{-3}, \mathrm{P}$ (Mehlich): 32.8 mg $\mathrm{dm}^{-3}, \mathrm{~K}: 0.532 \mathrm{cmol}_{\mathrm{c}} \mathrm{dm}^{-3}, \mathrm{CEC}_{\mathrm{pH7}}: 11.6 \mathrm{cmol}_{\mathrm{c}}$ $\mathrm{dm}^{-3}, \mathrm{Cu}: 2.2 \mathrm{mg} \mathrm{dm}^{-3}, \mathrm{Zn}: 1.01 \mathrm{mg} \mathrm{dm}^{-3}$, and B: $0.2 \mathrm{mg} \mathrm{dm}^{-3}$. These results were used to define the type of fertilisation (Comissão de Química e Fertilidade do Solo [CQFS], 2016).

On each of the following days (18/12/2017, 03/01/2018, 14/03/2018, 06/11/2018, 28/12/2018, 30/01/2019, 22/02/2019 and 28/03/2019), two uniformity trials were set up, one with the IPR91-Baili cultivar and the other with the IPR92-Altar cultivar, giving a total of 16 uniformity trials. In each trial, sowing was carried out in rows spaced 0.5 m apart at a density of $50 \mathrm{~kg} \mathrm{ha}^{-1}$ and with a base fertilisation of $35 \mathrm{~kg} \mathrm{ha}^{-1} \mathrm{~N}$, $135 \mathrm{~kg} \mathrm{ha}^{-1} \mathrm{P}_{2} \mathrm{O}_{5}$ and $135 \mathrm{~kg} \mathrm{ha}^{-1} \mathrm{~K}_{2} \mathrm{O}$. The fresh matter (FM) was evaluated on 15/02/2018, 14/03/2018, 10/05/2018, 27/12/2018, 14/02/2019, 03/04/2019, 15/04/2019 and 20/05/2019 respectively, i.e. 59, 70, 57, 51, 48, 63,52 and 53 days after sowing, to coincide with the final flowering period and the start of grain formation.

During the uniformity trials, the daily data for minimum (Tmin) and maximum (Tmax) air temperature, in ${ }^{\circ} \mathrm{C}$; insolation, in hours day ${ }^{-1}$; solar radiation, in $\mathrm{MJ} \mathrm{m}^{-2}$ day $^{-1}$; and rainfall, in
mm were recorded by the Weather Station of the 8th Meteorological District of the National Institute of Meteorology, located 100 m from the experimental area. The average daily air temperature (Tave), in ${ }^{\circ} \mathrm{C}$, was calculated from the expression: Tave $=(\operatorname{Tmin}+\operatorname{Tmax}) / 2$.

For each uniformity trial measuring $8 \mathrm{~m} \times 8 \mathrm{~m}\left(64 \mathrm{~m}^{2}\right)$, the central area, $6 \mathrm{~m} \times 6$ $\mathrm{m}\left(36 \mathrm{~m}^{2}\right)$ in size was divided into 36 basic experimental units (BEU), each of $1 \mathrm{~m} \times 1 \mathrm{~m}$ ( $1 \mathrm{~m}^{2}$ ), forming a matrix of six rows and six columns (Figure 1). In each BEU, the plants were cut close to the ground, and the FM was immediately determined, in $\mathrm{g} \mathrm{m}^{-2}$, on a digital scale (accuracy: 1 g ). Weighing was carried out immediately after cutting in order to minimise possible variations in plant moisture. Samples of fresh matter were collected and weighed (FMS, in g sample ${ }^{-1}$ ) from six randomly chosen BEU in each uniformity trial. The samples were packed in paper packages, identified by BEU, and dried in a forced air ventilation oven at $65 \pm 3^{\circ} \mathrm{C}$ to constant weight, i.e. the dry matter weight (DMS, in g sample ${ }^{-1}$ ). The dry matter content, in \%, was calculated for each uniformity trial using the expression: $\mathrm{DM}(\%)=$ DMS/FMS $\times 100$.

For each uniformity trial, plots with $X_{R}$ adjacent BEUs on a row and $\mathrm{X}_{\mathrm{c}}$ adjacent BEUs in a column were projected, using the FM data of the 36 BEUs. Plots of different sizes and/or shapes were projected $\left(X=X_{R} \times X_{C}\right)$, i.e. $(1 \times 1)$, $(1 \times 2),(1 \times 3),(1 \times 6),(2 \times 1),(2 \times 2),(2 \times 3),(2 \times 6),(3 \times 1)$, $(3 \times 2),(3 \times 3),(3 \times 6),(6 \times 1)),(6 \times 2)$ and $(6 \times 3)$. The abbreviations $X_{R^{\prime}} X_{C}$ and $X$, refer, respectively, to the number of adjacent BEUs in a row, the number of adjacent BEUs in a column and the plot size in number of BEUs. For each plot size $(\mathrm{X})$, the following were determined: n - number of plots with a size of $X$ BEU $(n=36 / X), M_{(X)}-$ mean value for plots with a size of $X$ BEU, and $\mathrm{CV}_{(\mathrm{x})}$ - coefficient of variation (in \%) between plots with a size of $X$ BEU. In each of the 16 trials, the optimal plot size (Xo) was determined using the methods of modified maximum curvature (MMC) (Meier \& Lessman, 1971), the linear response plateau model (LRP) (Paranaíba et al., 2009a) and the quadratic response plateau model (QRP) (Peixoto et al., 2011). In these three methods, models of the dependent variable ( $\mathrm{CV}_{(\mathrm{X})}$, in \%) are adjusted as a function of the independent variable ( $X$, in BEU).


Figure 1. Sketch of one uniformity trial measuring $8 \mathrm{~m} \times 8 \mathrm{~m}\left(64 \mathrm{~m}^{2}\right)$ with the central area of $6 \mathrm{~m} \times 6 \mathrm{~m}\left(36 \mathrm{~m}^{2}\right)$ divided into 36 basic experimental units (BEU) of $1 \mathrm{~m} \times 1 \mathrm{~m}\left(1 \mathrm{~m}^{2}\right)$, forming a matrix of six rows and six columns.

With reference to the method of modified maximum curvature (MMC) (Meier \& Lessman, 1971), the parameters a and b were estimated, together with the coefficient of determination ( $\mathrm{R}^{2}$ ) of the model $\mathrm{CV}_{(\mathrm{x})}=\mathrm{a} / \mathrm{X}^{\mathrm{b}}+\varepsilon$. These parameters were estimated through the logarithmic transformation and linearisation of $\mathrm{CV}_{(x)}=\mathrm{a} / \mathrm{X}^{\mathrm{b}}+\varepsilon$, i.e. $\log \mathrm{CV}_{(\mathrm{x})}=\log \mathrm{a}-\mathrm{b} \log$ $X+\varepsilon$, where the estimate was weighted by the degrees of freedom (DF = $n-1$ ) associated with each plot size (Steel, Torrie, \& Dickey, 1997). The point corresponding to the optimal plot size ( Xo ) was algebraically determined using the expression: $X o=\left[a^{2} b^{2}(2 b+1) /(b+2)\right]^{1 /(2 b+2)}$. The coefficient of variation corresponding to the optimal plot size $\left(\mathrm{CV}_{\mathrm{xo}_{0}}\right)$ was determined by $C_{X_{0}}=a / X o^{b}$.

For the linear response plateau model (LRP) (Paranaíba et al., 2009a), two segmented lines were adjusted and estimates of the parameters $a, b$ and $p$, and the coefficient of
determination ( $\mathrm{R}^{2}$ ) were obtained. The first line $\left(\mathrm{CV}_{(\mathrm{X})}=a+b X+\varepsilon\right)$ is adjusted to the point corresponding to the optimal plot size ( Xo ) with a non-zero slope (b). The second line ( $\mathrm{CV}_{(x)}=p+$ $\varepsilon)$ starts at Xo and has zero slope, i.e. it is a line parallel to the abscissa, where $p=$ the plateau, i.e. p corresponds to $\mathrm{CV}_{\text {xo }^{\prime}}$. The LRP model was as follows:

$$
\mathrm{CV}_{(X)}= \begin{cases}a+b X+\varepsilon & \text { if } X \leq X o \\ p+\varepsilon & \text { if } X>X o .\end{cases}
$$

In the LRP model, the optimal plot size was determined by $X o=(p-a) / b$, and the coefficient of variation at the optimal plot size by $\mathrm{CV}_{\mathrm{xo}_{0}}=a$ $+b X o$.

For the quadratic response plateau model (QRP) (Peixoto etal., 2011), the adjustment was made using two segmented equations. Estimates of parameters $a, b, c$ and $p$, and of the coefficient of determination ( $\mathrm{R}^{2}$ ) were obtained. The quadratic part of the model was adjusted up to point $X o\left(C V_{(x)}=a+b X+c X^{2}+\varepsilon\right)$. From Xo the
model becomes a straight line with zero slope, called the plateau, whose model is described by $\left(C V_{(x)}=p+\varepsilon\right)$ where $p=$ the plateau, i.e. $p=C V_{x_{0}}$. Therefore, the QRP model was as follows:

$$
C V_{(x)}= \begin{cases}a+b X+c X^{2}+\varepsilon & \text { if } X \leq X o \\ p+\varepsilon & \text { if } X>X o .\end{cases}
$$

In the QRP model, the optimal plot size was determined by $X o=-b / 2 c$ and the coefficient of variation at the optimal plot size by $C V_{x_{0}}=a-$ $b^{2 / 4 c}$. In the LRP and QRP models, the point of union between the two segments corresponds to Xo on the abscissa and $\mathrm{CV}_{\mathrm{x}_{0}}$ on the ordinate. In the three models (MMC, LRP and QRP), $\varepsilon$ represents the residual or random error of the model.

Therefore, for each uniformity trial within each cultivar, estimates of the coefficient of determination $\left(\mathrm{R}^{2}\right)$, the optimal plot size (Xo) and the coefficient of variation at the optimal plot size $\left(\mathrm{CV}_{\mathrm{xo}_{0}}\right)$ were obtained relative to the MMC, LRP and QRP methods. The mean values of the estimates for $\mathrm{R}^{2}$, Xo and $\mathrm{CV}_{\mathrm{xo}^{\prime}}$, in relation to the MMC, LRP and QRP methods, were compared between the buckwheat cultivars (IPR91-Baili versus IPR92-Altar) within each method ( $\mathrm{n}=8$ uniformity trials), using Student's t-test (unilateral) for independent samples, at a significance level of 5\%. Comparisons were made between the methods (MMC versus LRP, MMC versus QRP and LRP versus QRP) within each buckwheat cultivar ( $\mathrm{n}=8$ uniformity trials) and overall, i.e. irrespective of the cultivar ( $\mathrm{n}=16$ uniformity trials), using Student's t-test (unilateral) for dependent samples, at a significance level of $5 \%$. The results of these comparisons were represented by letters at the side of each mean value. The statistical analysis
was carried out with the aid of Microsoft Office Excel ${ }^{\circledR}$ and the R ( R Development Core Team [R], 2020) and Genes (Cruz, 2016) software.

## Results and Discussion

Among the uniformity trials, the fresh matter (FM) of the IPR91-Baili buckwheat cultivar ranged from 459 to $3336 \mathrm{~g} \mathrm{~m}^{-2}(4.59$ and $33.36 \mathrm{Mg} \mathrm{ha}^{-1}$ ) with a mean value of 1870 $\mathrm{g} \mathrm{m}^{-2}\left(18.70 \mathrm{Mg}_{\mathrm{ha}}{ }^{-1}\right)$, and from 474 to 3311 $\mathrm{g} \mathrm{m}^{-2}$ ( 4.74 and $33.11 \mathrm{Mg} \mathrm{ha}^{-1}$ ) with a mean value of $1833 \mathrm{~g} \mathrm{~m}^{-2}$ ( $18.33 \mathrm{Mg}^{\mathrm{Ma}}{ }^{-1}$ ) for the IPR92-Altar cultivar (Tables 1 and 2). The difference between the mean values for FM in the cultivars was not significant ( $\mathrm{t}=0.068$; p-value $=0.9467 ; 14$ degrees of freedom). The dry matter weight (DM) obtained for the sowing times of 18/12/2017, 03/01/2018, 14/03/2018, 06/11/2018, 28/12/2018, 30/01/2019, 22/02/2019, 28/03/2019 were 505, 591, 134, 177, 259, 515, 533 and 107 g $\mathrm{m}^{-2}$ respectively, for the IPR91-Baili cultivar, and $490,557,174,202,99,482,494$ and $92 \mathrm{~g} \mathrm{~m}^{-2}$, for the IPR92-Altar cultivar. Therefore, the DM ranged between 1.07 and $5.91 \mathrm{Mg} \mathrm{ha}^{-1}$, with a mean value of $3.53 \mathrm{Mg}_{\mathrm{ha}}{ }^{-1}$ for the IPR91-Baili cultivar and between 0.92 and $5.57 \mathrm{Mg} \mathrm{ha}^{-1}$, with a mean of $3.24 \mathrm{Mg} \mathrm{ha}^{-1}$ for the IPR92Altar cultivar; the difference between mean values was not significant ( $\mathrm{t}=0.286$; p -value $=0.7787 ; 14$ degrees of freedom). As such, it can be inferred that between the IPR91-Baili and IPR92-Altar cultivars there were similar amounts of FM and DM.

Table 1
Planned plot size ( $\mathrm{X}=\mathrm{X}_{\mathrm{R}} \times \mathrm{X}_{\mathrm{C}}$ ), in basic experimental units (BEU), with $\mathrm{X}_{\mathrm{R}}$ adjacent BEUs in a row and $\mathrm{X}_{\mathrm{c}}$ adjacent BEUs in a column; number of plots with a size of $X$ BEU ( $n=36 / X$ ); mean value for plots with a size of $X$ BEU $\left[M_{(x)}\right]$, in g ; and coefficient of variation (in \%) between plots with a size of X BEU [CV $\left.{ }_{(x)}\right]$. Fresh matter data in buckwheat (Fagopyrum esculentum Moench) 'IPR91-Baili', obtained in uniformity trials conducted at eight sowing times

| X | $\mathrm{X}_{\mathrm{c}}$ | X | n | $\mathrm{M}_{(x)}$ | CV ${ }_{(x)}$ | $M_{(x)}$ | $\mathrm{CV}_{(x)}$ | $\mathrm{M}_{(x)}$ | CV ${ }_{(x)}$ | $\mathrm{M}_{(x)}$ | $\mathrm{CV}_{(x)}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | 18/12/2017 |  | 03/01/2018 |  | 14/03/2018 |  | 06/11/2018 |  |
| 1 | 1 | 1 | 36 | 3336 | 19.37 | 3105 | 14.19 | 901 | 26.84 | 1227 | 30.53 |
| 1 | 2 | 2 | 18 | 6673 | 12.69 | 6210 | 12.54 | 1802 | 22.52 | 2455 | 23.75 |
| 1 | 3 | 3 | 12 | 10009 | 9.03 | 9314 | 11.01 | 2703 | 17.26 | 3682 | 18.98 |
| 1 | 6 | 6 | 6 | 20019 | 7.01 | 18629 | 10.38 | 5405 | 7.69 | 7365 | 16.28 |
| 2 | 1 | 2 | 18 | 6673 | 14.59 | 6210 | 10.85 | 1802 | 23.29 | 2455 | 26.42 |
| 2 | 2 | 4 | 9 | 13346 | 9.07 | 12419 | 9.20 | 3604 | 21.31 | 4910 | 19.91 |
| 2 | 3 | 6 | 6 | 20019 | 5.94 | 18629 | 7.86 | 5405 | 15.46 | 7365 | 14.86 |
| 2 | 6 | 12 | 3 | 40038 | 1.88 | 37258 | 7.66 | 10811 | 7.43 | 14729 | 13.37 |
| 3 | 1 | 3 | 12 | 10009 | 15.21 | 9314 | 8.15 | 2703 | 23.37 | 3682 | 24.47 |
| 3 | 2 | 6 | 6 | 20019 | 9.30 | 18629 | 7.11 | 5405 | 22.07 | 7365 | 17.46 |
| 3 | 3 | 9 | 4 | 30028 | 7.09 | 27943 | 5.50 | 8108 | 15.85 | 11047 | 11.74 |
| 3 | 6 | 18 | 2 | 60057 | 6.08 | 55887 | 6.10 | 16216 | 5.15 | 22094 | 10.00 |
| 6 | 1 | 6 | 6 | 20019 | 10.48 | 18629 | 6.33 | 5405 | 22.84 | 7365 | 24.10 |
| 6 | 2 | 12 | 3 | 40038 | 7.71 | 37258 | 4.86 | 10811 | 22.90 | 14729 | 17.35 |
| 6 | 3 | 18 | 2 | 60057 | 5.37 | 55887 | 2.85 | 16216 | 16.29 | 22094 | 10.22 |


|  |  |  |  | 28/12/2018 |  | 30/01/2019 |  | 22/02/2019 |  | 28/03/2019 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 1 | 36 | 1397 | 17.17 | 2470 | 23.17 | 2067 | 16.30 | 459 | 41.40 |
| 1 | 2 | 2 | 18 | 2795 | 11.98 | 4940 | 15.26 | 4135 | 13.96 | 918 | 36.18 |
| 1 | 3 | 3 | 12 | 4192 | 10.41 | 7409 | 11.08 | 6202 | 11.29 | 1377 | 31.49 |
| 1 | 6 | 6 | 6 | 8384 | 7.65 | 14819 | 9.87 | 12405 | 7.67 | 2753 | 22.15 |
| 2 | 1 | 2 | 18 | 2795 | 14.04 | 4940 | 15.90 | 4135 | 11.50 | 918 | 24.69 |
| 2 | 2 | 4 | 9 | 5589 | 8.90 | 9879 | 10.38 | 8270 | 9.39 | 1836 | 22.13 |
| 2 | 3 | 6 | 6 | 8384 | 7.94 | 14819 | 6.84 | 12405 | 9.69 | 2753 | 20.11 |
| 2 | 6 | 12 | 3 | 16767 | 4.18 | 29638 | 5.95 | 24809 | 5.00 | 5507 | 21.34 |
| 3 | 1 | 3 | 12 | 4192 | 11.32 | 7409 | 13.83 | 6202 | 10.43 | 1377 | 22.75 |
| 3 | 2 | 6 | 6 | 8384 | 7.03 | 14819 | 8.60 | 12405 | 8.87 | 2753 | 21.38 |
| 3 | 3 | 9 | 4 | 12575 | 6.88 | 22228 | 3.23 | 18607 | 9.05 | 4130 | 19.02 |
| 3 | 6 | 18 | 2 | 25151 | 5.16 | 44457 | 1.49 | 37214 | 4.87 | 8260 | 23.03 |
| 6 | 1 | 6 | 6 | 8384 | 9.48 | 14819 | 13.66 | 12405 | 9.84 | 2753 | 15.13 |
| 6 | 2 | 12 | 3 | 16767 | 5.12 | 29638 | 9.33 | 24809 | 8.90 | 5507 | 12.77 |
| 6 | 3 | 18 | 2 | 25151 | 6.64 | 44457 | 3.66 | 37214 | 9.86 | 8260 | 3.51 |

(1) Each uniformity trial of $6 \mathrm{~m} \times 6 \mathrm{~m}(36 \mathrm{~m} 2)$ was divided into 36 BEU of $1 \mathrm{~m} \times 1 \mathrm{~m}\left(1 \mathrm{~m}^{2}\right)$, forming a matrix of six rows and six columns.

The FM evaluated at 90 days after plant emergence in buckwheat by Menezes and Leandro (2004) was $3.58 \mathrm{Mg} \mathrm{ha}^{-1}$. At 93 days after sowing, i.e. at the full flowering stage, Ziech et al. (2015) obtained a DM of $2.8 \mathrm{Mg} \mathrm{ha}^{-1}$. In cuts made at 47,57 and 67 days after sowing buckwheat, Görgen et al. (2016) obtained DM of $2.301,3.144$ and $4.471 \mathrm{Mg}^{\text {ha }}{ }^{-1}$ respectively. At 71 days after sowing buckwheat, during the reproductive period between flowering and the grain milk stage, Pereira et al. (2017) obtained values for FM and DM of 26.97
and $6.78 \mathrm{Mg} \mathrm{ha}^{-1}$ respectively. Fresh matter weights in buckwheat of less than 3.00 Mg ha ${ }^{-1}$ were obtained by Skora and Campos (2017). The different environmental conditions, management, cultivars and evaluation times make comparisons difficult. Even so, in general, it can be seen that the values obtained in this study were similar to those reported in the above studies, and demonstrated good plant development under the environmental conditions of the area (Table 3).

Table 2
Planned plot size ( $X=X_{R} \times X_{C}$ ), in basic experimental units (BEU), with $X_{R}$ adjacent BEUs in a row and $X_{c}$ adjacent BEUs in a column; number of plots with a sized of $X$ BEU ( $n=36 / X$ ); mean value for plots with a size of $X$ BEU $\left[M_{(x)}\right]$, in $g$; and coefficient of variation (in \%) between plots with a size of X BEU [CV ${ }_{(x)}$ ]. Fresh matter data in buckwheat (Fagopyrum esculentum Moench) 'IPR92-Altar', obtained in uniformity trials conducted at eight sowing times

| X | $\mathrm{X}_{\mathrm{c}}$ | X | n | $\mathrm{M}_{(1)}$ | CV ${ }_{(x)}$ | $\mathrm{M}_{(\times)}$ | $\mathrm{CV}_{(x)}$ | $\mathrm{M}_{(\times)}$ | CV ${ }_{(x)}$ | $\mathrm{M}_{(x)} \quad \mathrm{CV}_{(x)}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | 18/12/2017 |  | 03/01/2018 |  | 14/03/2018 |  | 06/11/2018 |  |
| 1 | 1 | 1 | 36 | 3253 | 17.53 | 3311 | 21.26 | 1009 | 23.41 | 1245 | 33.59 |
| 1 | 2 | 2 | 18 | 6505 | 13.14 | 6621 | 13.26 | 2019 | 20.89 | 2490 | 29.37 |
| 1 | 3 | 3 | 12 | 9758 | 9.76 | 9932 | 10.67 | 3028 | 19.67 | 3735 | 24.49 |
| 1 | 6 | 6 | 6 | 19516 | 4.78 | 19863 | 8.20 | 6056 | 14.44 | 7471 | 17.63 |
| 2 | 1 | 2 | 18 | 6505 | 11.82 | 6621 | 15.49 | 2019 | 15.02 | 2490 | 27.69 |
| 2 | 2 | 4 | 9 | 13011 | 10.68 | 13242 | 7.92 | 4038 | 13.84 | 4981 | 25.49 |
| 2 | 3 | 6 | 6 | 19516 | 4.85 | 19863 | 3.77 | 6056 | 13.40 | 7471 | 20.85 |
| 2 | 6 | 12 | 3 | 39032 | 5.17 | 39726 | 2.51 | 12113 | 4.79 | 14942 | 11.79 |
| 3 | 1 | 3 | 12 | 9758 | 11.48 | 9932 | 14.98 | 3028 | 14.69 | 3735 | 22.67 |
| 3 | 2 | 6 | 6 | 19516 | 10.45 | 19863 | 7.50 | 6056 | 13.28 | 7471 | 20.88 |
| 3 | 3 | 9 | 4 | 29274 | 4.14 | 29795 | 2.62 | 9085 | 13.56 | 11206 | 15.49 |
| 3 | 6 | 18 | 2 | 58548 | 4.79 | 59589 | 0.86 | 18169 | 8.87 | 22413 | 11.23 |
| 6 | 1 | 6 | 6 | 19516 | 10.24 | 19863 | 14.25 | 6056 | 12.90 | 7471 | 16.11 |
| 6 | 2 | 12 | 3 | 39032 | 10.22 | 39726 | 7.54 | 12113 | 12.51 | 14942 | 17.29 |
| 6 | 3 | 18 | 2 | 58548 | 1.56 | 59589 | 3.07 | 18169 | 14.01 | 22413 | 14.71 |


|  |  |  |  | 28/12/2018 |  | 30/01/2019 |  | 22/02/2019 |  | 28/03/2019 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 1 | 36 | 621 | 32.57 | 2635 | 24.81 | 2117 | 20.97 | 474 | 29.96 |
| 1 | 2 | 2 | 18 | 1242 | 20.93 | 5269 | 15.82 | 4234 | 17.20 | 948 | 24.11 |
| 1 | 3 | 3 | 12 | 1863 | 22.05 | 7904 | 15.10 | 6350 | 14.45 | 1422 | 21.61 |
| 1 | 6 | 6 | 6 | 3727 | 8.33 | 15807 | 9.65 | 12701 | 13.08 | 2844 | 12.72 |
| 2 | 1 | 2 | 18 | 1242 | 26.24 | 5269 | 15.54 | 4234 | 16.81 | 948 | 23.98 |
| 2 | 2 | 4 | 9 | 2484 | 16.72 | 10538 | 7.56 | 8467 | 14.71 | 1896 | 18.97 |
| 2 | 3 | 6 | 6 | 3727 | 20.10 | 15807 | 8.07 | 12701 | 11.69 | 2844 | 19.71 |
| 2 | 6 | 12 | 3 | 7453 | 7.40 | 31614 | 3.21 | 25401 | 10.33 | 5688 | 11.39 |
| 3 | 1 | 3 | 12 | 1863 | 25.78 | 7904 | 14.81 | 6350 | 11.00 | 1422 | 21.89 |
| 3 | 2 | 6 | 6 | 3727 | 17.96 | 15807 | 6.66 | 12701 | 7.89 | 2844 | 18.07 |
| 3 | 3 | 9 | 4 | 5590 | 21.62 | 23711 | 4.22 | 19051 | 5.97 | 4266 | 19.65 |
| 3 | 6 | 18 | 2 | 11180 | 10.27 | 47421 | 0.05 | 38102 | 3.12 | 8533 | 12.54 |
| 6 | 1 | 6 | 6 | 3727 | 23.03 | 15807 | 8.95 | 12701 | 8.64 | 2844 | 17.57 |
| 6 | 2 | 12 | 3 | 7453 | 16.88 | 31614 | 4.86 | 25401 | 6.62 | 5688 | 15.00 |
| 6 | 3 | 18 | 2 | 11180 | 24.38 | 47421 | 0.18 | 38102 | 3.70 | 8533 | 16.85 |

[^1]Table 3
Monthly mean value for the minimum (Tmin), maximum (Tmax) and average (Tave) air temperature, in ${ }^{\circ} \mathrm{C}$; monthly mean value for daily insolation, in hrs day ${ }^{-1}$; monthly mean value for daily solar radiation, in MJ m${ }^{-2}$ day $^{-1}$; total rainfall, in mm, for the evaluation periods of buckwheat (Fagopyrum esculentum Moench) 'IPR91-Baili' and 'IPR92-Altar', and the climate normal for 1981-2010
$\left.\begin{array}{|cccccccc|}\hline \text { Month } & \text { Year } & \begin{array}{c}\text { Tmin } \\ \left({ }^{\circ} \mathrm{C}\right)\end{array} & \begin{array}{c}\text { Tmax } \\ \left({ }^{\circ} \mathrm{C}\right)\end{array} & \begin{array}{c}\text { Tave } \\ \left({ }^{\circ} \mathrm{C}\right)\end{array} & \begin{array}{c}\text { Insolation } \\ \left(\mathrm{hrs} \mathrm{day}{ }^{-1}\right)\end{array} & \begin{array}{c}\text { Radiation } \\ \left(\mathrm{MJ} \mathrm{m}^{-2} \text { day }{ }^{-1}\right)\end{array} & \begin{array}{c}\text { Rainfall } \\ (\mathrm{mm})\end{array} \\ \hline \text { November } & 2017 / 2018 & \text { agricultural year (sown on } & 18 / 12 / 2017,03 / 01 / 2018 & \text { and } 14 / 03 / 2018)\end{array}\right]$

2018/2019 agricultural year (sown on 06/11/2018, 28/12/2018, 30/01/2019, 22/02/2019 and 28/03/2019)

| November | 2018 | 17.7 | 28.8 | 23.3 | 7.9 | 22.8 | 245.1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| December | 2018 | 18.7 | 30.0 | 24.4 | 8.3 | 24.4 | 176.2 |
| January | 2019 | 21.9 | 30.8 | 26.4 | 5.3 | 19.0 | 266.1 |
| February | 2019 | 19.1 | 30.0 | 24.6 | 8.1 | 21.7 | 83.4 |
| March | 2019 | 17.2 | 28.1 | 22.7 | 7.2 | 18.3 | 136.3 |
| April | 2019 | 17.3 | 26.4 | 21.8 | 5.1 | 12.6 | 210.4 |
| May | 2019 | 15.4 | 22.5 | 19.0 | 3.6 | 8.2 | 260.2 |


|  | Climate normal 1981-2010 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| November | 16.4 | 27.8 | 21.6 | 7.3 | - | 132.7 |
| December | 18.6 | 30.3 | 24.1 | 8.1 | - | 154.3 |
| January | 19.8 | 30.9 | 24.9 | 7.9 | - | 166.3 |
| February | 19.4 | 29.9 | 24.0 | 7.1 | - | 139.6 |
| March | 18.5 | 29.1 | 22.9 | 6.8 | - | 127.7 |
| April | 15.0 | 25.7 | 19.4 | 5.8 | - | 170.1 |
| May | 12.0 | 22.0 | 16.0 | 5.0 | - | 154.4 |

From the FM data of the 36 basic experimental units (BEU) of $1 \mathrm{~m} \times 1 \mathrm{~m}\left(1 \mathrm{~m}^{2}\right)$, the coefficient of variation (CV) varied between 14.19 and $41.40 \%$ for the IPR91-Baili cultivar and between 17.53 and $33.59 \%$ for the IPR92Altar cultivar (Tables 1 and 2). This variations in the CV may be associated with environmental and genotypic variability, and the genotype interaction with the environment. The mean coefficient of variation was 23.62\% and 25.51\% respectively, for the IPR91-Baili and IPR92-Altar cultivars, and the difference between the mean values was not significant ( $t=-0.494 ; p$-value = 0.6287; 14 degrees of freedom). According to the classification bands of the coefficients of variation established by Pimentel-Gomes (2009) for agricultural field trials, the values obtained are classified as of medium (between 10 and 20\%), low (between 20 and 30\%) and very low (>30\%) experimental precision, which shows that it is necessary to use a plot size greater than $1 \mathrm{~m}^{2}$ to improve the experimental precision.

This wide variability in FM and CV between the uniformity trials is possibly associated with the different environmental conditions between the years and sowing times (Table 3). Furthermore, this scenario of wide variability gives credibility to the study of plot size, since it includes real situations of variability that occur in experiments conducted in the field. The non-significant difference between cultivars, in relation to FM and CV, suggests that the plot size for experiments with these two buckwheat cultivars may be similar.

In each uniformity trial conducted with the IPR91-Baili and IPR92-Altar cultivars, there was an increase in the mean value of the plots $\left[\mathrm{M}_{(x)}\right]$ and a reduction in the coefficient of variation $\left[\mathrm{CV}_{(x)}\right]$ for an increase in the planned plot size (X) (Tables 1 and 2). These results indicate an improvement in experimental precision (a reduction in $\mathrm{CV}_{(x)}$ ) with the increase in plot size. Although it is possible to evaluate fresh matter (FM) in plots of $1 \mathrm{~m}^{2}$, based on the methodology used in this study, it is important to evaluate the precision of larger-sized plots, i.e. it is essential to plan the experiment at the optimal plot size to ensure proper discrimination of the treatments under evaluation and the reliability of the inferences. In addition, smaller sizes may not represent plant development, while larger sizes make it possible to evaluate the plants in the central area of each plot (working area) and disregard the borders, thereby reducing interference from the plants of adjacent plots, e.g. inter-plot competition (Storck et al., 2016).

For the methods of modified maximum curvature (MMC), the linear response plateau model (LRP) and the quadratic response plateau model (QRP), the mean values for the coefficient of determination ( $\mathrm{R}^{2}$ ), the optimal plot size (Xo) and the coefficient of variation at the optimal plot size $\left(\mathrm{CV}_{\mathrm{xo}_{0}}\right)$ did not differ between the two cultivars (IPR91-Baili and IPR92-Altar) (Tables 4 and 5). As such, based on these results and on the lack of difference in FM, DM and CV between the cultivars, it can be inferred that the experimental planning of plot size is similar for both cultivars.

Table 4
Estimates of the coefficient of determination ( $\mathrm{R}^{2}$ ), optimal plot size ( Xo , in $\mathrm{m}^{2}$ ) and coefficient of variation at the optimal plot size ( $C V_{\mathrm{Xo}^{\prime}}$ in \%), in relation to the methods of modified maximum curvature (MMC), the linear response plateau model (LRP) and the quadratic response plateau model (QRP), obtained from the fresh matter of buckwheat (Fagopyrum esculentum Moench) 'IPR91-Baili' and 'IPR92-Altar', in eight uniformity trials

| Trial | Sown | MMC |  |  | LRP |  |  | QRP |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\mathrm{R}^{2}$ | Xo | $\mathrm{CV}_{\text {xo }}$ | $\mathrm{R}^{2}$ | Xo | $\mathrm{CV}_{\text {xo }}$ | $\mathrm{R}^{2}$ | Хо | $\mathrm{CV}_{\text {¢ }}$ |
| Cultivar IPR91-Baili ( $\mathrm{n}=8$ uniformity trials) |  |  |  |  |  |  |  |  |  |  |
| 1 | 18/12/2017 | 0.78 | 4.28 | 9.27 | 0.74 | 7.17 | 5.63 | 0.77 | 9.31 | 5.73 |
| 2 | 03/01/2018 | 0.82 | 3.03 | 9.69 | 0.76 | 8.21 | 5.39 | 0.77 | 11.51 | 5.39 |
| 3 | 14/03/2018 | 0.48 | 4.39 | 17.82 | 0.39 | 8.07 | 13.52 | 0.41 | 19.06 | 11.33 |
| 4 | 06/11/2018 | 0.83 | 5.03 | 18.26 | 0.77 | 8.57 | 12.54 | 0.77 | 11.68 | 12.66 |
| 5 | 28/12/2018 | 0.94 | 3.76 | 9.65 | 0.88 | 7.35 | 5.60 | 0.91 | 9.99 | 5.60 |
| 6 | 30/01/2019 | 0.83 | 4.96 | 9.17 | 0.76 | 8.20 | 4.73 | 0.78 | 10.52 | 5.03 |
| 7 | 22/02/2019 | 0.85 | 3.08 | 11.05 | 0.67 | 6.94 | 7.54 | 0.71 | 8.05 | 7.75 |
| 8 | 28/03/2019 | 0.73 | 6.73 | 18.56 | 0.61 | 6.83 | 15.93 | 0.64 | 8.29 | 16.32 |


| Mínimum | 0.48 | 3.03 | 9.17 | 0.39 | 6.83 | 4.73 | 0.41 | 8.05 | 5.03 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Maximum | 0.94 | 6.73 | 18.56 | 0.88 | 8.57 | 15.93 | 0.91 | 19.06 | 16.32 |
| $\quad$ Mean | 0.78 | 4.41 | 12.93 | 0.70 | 7.67 | 8.86 | 0.72 | 11.05 | 8.73 |


| 1 | $18 / 12 / 2017$ | 0.75 | 3.99 | 8.91 | 0.67 | 7.41 | 5.17 | 0.69 | 9.72 | 5.30 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | $03 / 01 / 2018$ | 0.76 | 4.93 | 7.63 | 0.74 | 8.23 | 3.32 | 0.76 | 11.03 | 3.48 |
| 3 | $14 / 03 / 2018$ | 0.70 | 3.98 | 14.82 | 0.62 | 7.51 | 10.75 | 0.64 | 11.76 | 10.38 |
| 4 | $06 / 11 / 2018$ | 0.91 | 5.50 | 19.69 | 0.89 | 7.81 | 14.10 | 0.90 | 11.46 | 13.87 |
| 5 | $28 / 12 / 2018$ | 0.58 | 5.24 | 17.83 | 0.40 | 6.20 | 16.11 | 0.44 | 5.75 | 16.65 |
| 6 | $30 / 01 / 2019$ | 0.68 | 5.56 | 6.48 | 0.87 | 7.96 | 2.50 | 0.90 | 11.43 | 2.33 |
| 7 | $22 / 02 / 2019$ | 0.82 | 4.47 | 11.14 | 0.80 | 8.37 | 5.95 | 0.81 | 16.88 | 4.39 |
| 8 | $28 / 03 / 2019$ | 0.87 | 4.66 | 18.82 | 0.75 | 6.68 | 15.08 | 0.78 | 8.11 | 15.28 |


| Mínimum | 0.58 | 3.98 | 6.48 | 0.40 | 6.20 | 2.50 | 0.44 | 5.75 | 2.33 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Maximum | 0.91 | 5.56 | 19.69 | 0.89 | 8.37 | 16.11 | 0.90 | 16.88 | 16.65 |
| Mean | 0.76 | 4.79 | 13.16 | 0.72 | 7.52 | 9.12 | 0.74 | 10.77 | 8.96 |
|  | 0.48 | 3.03 | 6.48 | 0.39 | 6.20 | 2.50 | 0.41 | 5.75 | 2.33 |
| Mínimum | 0.94 | 6.73 | 19.69 | 0.89 | 8.57 | 16.11 | 0.91 | 19.06 | 16.65 |
| Maximum | 0.77 | 4.60 | 13.05 | 0.71 | 7.60 | 8.99 | 0.73 | 10.91 | 8.84 |
| Mean |  |  |  |  |  | uniformity trials) |  |  |  |

Table 5
Mean estimates of the coefficient of determination ( $\mathrm{R}^{2}$ ), optimal plot size ( Xo , in $\mathrm{m}^{2}$ ) and coefficient of variation at the optimal plot size ( $\mathrm{CV}_{\mathrm{xo}^{\prime}}$ in \%), in relation to the methods of modified maximum curvature (MMC), the linear response plateau model (LRP) and the quadratic response plateau model (QRP), obtained from the fresh matter of buckwheat (Fagopyrum esculentum Moench) 'IPR91-Baili' and 'IPR92-Altar', in eight uniformity trials

| Cultivar | MMC | LRP | QRP |
| :---: | :---: | :---: | :---: |
| Coefficient of determination ( $\mathrm{R}^{2}$ ) |  |  |  |
| IPR91-Baili | 0.78 A a | 0.70 A c | 0.72 A b |
| IPR92-Altar | 0.76 A ab | 0.72 A a | 0.74 A b |
| Overall | 0.77 a | 0.71 c | 0.73 b |
| Optimal plot size ( Xo , in $\mathrm{m}^{2}$ ) |  |  |  |
| IPR91-Baili | 4.41 A c | 7.67 A b | 11.05 A a |
| IPR92-Altar | 4.79 A c | 7.52 A b | 10.77 A a |
| Overall | 4.60 c | 7.60 b | 10.91 a |
| Coefficient of variation for optimal plot size ( $\mathrm{CV}_{\mathrm{Xo}^{\prime} \text {, in \%) }}$ |  |  |  |
| IPR91-Baili | 12.93 A a | 8.86 A b | 8.73 A b |
| IPR92-Altar | 13.16 A a | 9.12 Ab | 8.96 A b |
| Overall | 13.05 a | 8.99 b | 8.84 b |

* Mean values not followed by the same uppercase letter in a column (comparing cultivars within each method; $n=8$ uniformity trials) differ by Student's t-test (unilateral) for independent samples, at a level of $5 \%$ significance. Mean values not followed by the same lowercase letter on a line (comparing methods within each cultivar, $\mathrm{n}=8$ uniformity trials; and overall, i.e. irrespective of the cultivar, $\mathrm{n}=16$ uniformity trials) differ by Student's t -test (unilateral) for dependent samples, at a level of $5 \%$ significance.

The mean values for $R^{2}$ varied between 0.70 (LRP - 'IPR91-Baili') and 0.78 (MMC - 'IPR91Baili'). It should be considered that $0.00 \leq R 2$ $\leq 1.00$, interpreted as the closer $\mathrm{R}^{2}$ is to 1.00 the better the model fits the data. Generally, between the MMC, LRP and QRP methods, the MMC values for $\mathrm{R}^{2}$ were higher, the QRP values were intermediate and the LRP values were lower in the uniformity trials with the IPR91Baili and IPR92-Altar cultivars, and overall (irrespective of cultivar) (Table 5). Specifically, in comparing methods within the IPR92-Altar cultivar, the $R^{2}$ value for MMC did not differ from that of LRP or QRP, however the $\mathrm{R}^{2}$ for LRP and QRP differed from each other. This result can be explained by the dependent samples being
compared in pairs (MMC versus LRP, MMC versus QRP and LRP versus QRP). As such, it can be inferred that although the adjustments were different, each method showed $\mathrm{R}^{2}$ close to one ( $\mathrm{R}^{2} \geq 0.70$ ).

In relation to the IPR91-Baili cultivar, the optimal plot size (Xo) differed between the three methods, with $11.05 \mathrm{~m}^{2}$ for QRP, $7.67 \mathrm{~m}^{2}$ for LRP and $4.41 \mathrm{~m}^{2}$ for MMC. There was also a difference between methods for the IPR92Altar cultivar, of $10.77 \mathrm{~m}^{2}$ for QRP, $7.52 \mathrm{~m}^{2}$ for LRP and $4.79 \mathrm{~m}^{2}$ for MMC (Table 5). When comparing the cultivars within each method the differences were not significant. It can therefore be inferred that the plot size might be
the same for these cultivars and depends on the method of estimation.

The coefficient of variation at the optimal plot size ( $\mathrm{CV}_{\mathrm{Xo}^{\prime}}$ in \%) for both cultivars was higher with the MMC method compared to LRP and QRP, which did not differ from each other (Table 5). Between cultivars within each method, the differences in $\mathrm{CV}_{\mathrm{x}}$ were not significant. These results indicate better experimental precision when using plot sizes determined by the LRP and QRP methods compared to the MMC method, irrespective of the cultivar.

Regardless of the cultivar, the mean values for $\mathrm{R}^{2}$ differed between methods (MMC $=0.77 ; ~ L R P=0.71 ; ~ Q R P=0.73)$. The mean values for Xo decreased in the following order: QRP $=10.91 \mathrm{~m}^{2} ; ~ L R P=7.60 \mathrm{~m}^{2}$; and $\mathrm{MMC}=$ $4.60 \mathrm{~m}^{2}$, whereas $\mathrm{CV}_{\text {xo }}$ was higher with MMC (13.05\%), and there was no difference between LRP (8.99\%) and QRP (8.84\%). Therefore, despite the plot size being different between the LRP ( $7.60 \mathrm{~m}^{2}$ ) and QRP ( $10.91 \mathrm{~m}^{2}$ ) methods, they result in similar experimental precision, as $\mathrm{CV}_{\mathrm{xo}_{0}}$ did not differ. This lack of difference can be explained by the fact that, for a given plot size, any gains in precision (reduction in $\mathrm{CV}_{\mathrm{xo}_{0}}$ ) from additions to the area of the plot are not significant. As such, it can be inferred that plots of $7.60 \mathrm{~m}^{2}$ are suitable for experimental planning. This suggestion for plots of $7.60 \mathrm{~m}^{2}$ is supported by practical feasibility in the field, and stabilised precision starting from that size, and can be used as a reference when planning experiments with buckwheat. This plot size is slightly larger than the $4 \mathrm{~m}^{2}$ used by Pereira et al. (2017), and smaller than the $25 \mathrm{~m}^{2}$ used by Menezes and Leandro (2004), the $25 \mathrm{~m}^{2}$ used by Ziech et al. (2015), the $20 \mathrm{~m}^{2}$ used by Görgen et al. (2016) and the $24 \mathrm{~m}^{2}$ used by Skora and

Campos (2017) in research with buckwheat and other species of ground cover.

Higher estimates for $\mathrm{R}^{2}$, lower for Xo and higher for $C V_{x}$ were obtained with the MMC method compared to the LRP in rice (Paranaíba et al., 2009a), wheat and cassava (Paranaíba et al., 2009b) and papaya (Brito et al., 2012). In passion fruit, greater values for $R^{2}$ and Xo, and lower values for $\mathrm{CV}_{\text {xo }}$ were obtained with the QRP method compared to the LRP method (Peixoto et al., 2011). In general, the above research, comparing different methods to determine optimal plot size, found similar results to the those of present study.

## Conclusion

The optimal plot size differs between methods, and decreases in the following order: quadratic response plateau model, linear response plateau model and modified maximum curvature. The optimal plot size for evaluating fresh matter in buckwheat 'IPR91Baili' and 'IPR92-Altar' is $7.60 \mathrm{~m}^{2}$.

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[^1]:    (1) Each uniformity trial of $6 \mathrm{~m} \times 6 \mathrm{~m}\left(36 \mathrm{~m}^{2}\right)$ was divided into 36 BEU of $1 \mathrm{~m} \times 1 \mathrm{~m}\left(1 \mathrm{~m}^{2}\right)$, forming a matrix of six rows and six columns.

